

GENERALIZED MOMENTS FOR THE SEQUENTIAL LAW

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We investigate asymptotic behavior of the generalized moments $E(X_n^\mu \ell(X_n))$ for the sequential law under assumption that $\lim_n \frac{1}{n} E(X_n)$ exists, where ℓ is a slowly varying function and $\mu \in \mathbb{R}^+$.

1. INTRODUCTION

Let a set of random variables (X_n) be defined by

$$(1) \quad P\{X_n = k\} = p_{nk} > 0, \quad 1 \leq k \leq n; \quad \sum_{k \leq n} p_{nk} = 1.$$

Accordingly, the expectation is $E(X_n) := \sum_{k \leq n} kp_{nk}$ and moments of the m -th order are $E(X_n^m) := \sum_{k \leq n} k^m p_{nk}$.

In [3] we introduced a concept of the generalised moments $E(K_\rho(X_n))$, where $K_\rho(x) := x^\rho \ell(x)$, $x > 0$; $K_\rho(0) := 0$, is a regularly varying function of index $\rho \in \mathbb{R}$.

Thus

$$E(K_\rho(X_n)) := \sum_{k \leq n} k^\rho \ell(k) p_{nk}, \quad \rho \in \mathbb{R}^+.$$

In [3] and [4] we posed the following problem:

If $E(X_n) \rightarrow \infty$, give a characterization of probability laws with the property

$$(2) \quad E(K_\rho(X_n)) \sim c_\rho K_\rho(E(X_n)), \quad a < \rho < b, \quad (n \rightarrow \infty)$$

where c_ρ is a constant independent of n .

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That asymptotic behavior of ρ -th moment could depend only on the first moment is not so strange as it seems. For example, an elementary inequality shows that $E(X_n^\rho) \geq (E(X_n))^rho$ for each $\rho \geq 1$ and $E(X_n^\rho) \leq (E(X_n))^rho$ for $0 < \rho < 1$.

In cited papers we proved that the relation (2) actually takes place with $c_\rho = 1$, for generalised Binomial and POISSON laws.

1. PRELIMINARIES

Throughout the paper we have to deal with the KARAMATA's class K_ρ of regularly varying functions or sequences.

It is well known that $f \in K_\rho$ if it could be expressed in the form $f(x) := x^\rho \ell(x)$, $\rho \in \mathbb{R}$, where ρ is the index of regular variation and $\ell(x) \in K_0$ is so-called slowly varying function i.e. positive, continuous and satisfying

$$\forall t > 0, \ell(tx) \sim \ell(x) \quad (x \rightarrow \infty).$$

Some examples of slowly varying functions are:

$$1, \log^a x, \log^b(\log x), \exp\left(\frac{\log x}{\log \log x}\right), \exp(\log^c x); \quad a, b \in \mathbb{R}, 0 < c < 1.$$

The theory of regular variation is presented in [1] and [2]. We quote here some facts from [1]:

Lemma 1. *If $a(x) \sim b(x) \rightarrow \infty$ ($x \rightarrow \infty$), then $K_\rho(a(x)) \sim K_\rho(b(x))$ ($x \rightarrow \infty$).*

Lemma 2. *For $\mu > 0$ we have $\sup_{x \leq y} x^\mu \ell(x) \sim y^\mu \ell(y)$ ($y \rightarrow \infty$).*

Lemma 3. (KARAMATA's Theorem) *Let f be positive and locally bounded in $[a, \infty)$ and $\sigma > -\rho$. Then the following are equivalent*

$$(i) \quad f \in K_\rho; \quad (ii) \quad \frac{x^\sigma f(x)}{\int_a^x \frac{t^\sigma f(t)}{t} dt} \rightarrow \sigma + \rho \quad (x \rightarrow \infty).$$

Lemma 4. (VUILLEUMIER [5]) *If $\sum_{k \leq n} q_{nk} \rightarrow 1$ ($n \rightarrow \infty$) and there exists $\varepsilon > 0$ such that $\sum_{k \leq n} k^{-\varepsilon} |q_{nk}| = O(n^{-\varepsilon})$ then $\sum_{k \leq n} \ell(k) q_{nk} \sim \ell(n)$ ($n \rightarrow \infty$), for each $\ell \in K_0$.*

Now, we define the Sequential Law by the following: Let $\{p_k\}_{k=1}^\infty$ be a sequence of positive real numbers and put in (1) $P\{X_n = k\} = p_{nk}$, where

$$p_{nk} := \frac{p_k}{\sum_{m \leq n} p_m}, \quad k = 1, 2, \dots, n.$$

We study the asymptotic behavior of $E(K_\rho(X_n))$ with respect to $E(X_n)$ and show that (2) holds with $c_\rho \neq 1$.

3. THE RESULTS

Throughout the rest of the paper we suppose that $\lim_n \frac{1}{n} E(X_n)$ exists. Hence, if $\lim_n \frac{1}{n} E(X_n) = c$ it is evident that $c \in [0, 1]$, and we have to deal with three cases: $0 < c < 1$, $c = 0$, $c = 1$.

Proposition 1. For $0 < c < 1$ and $\mu > -\frac{c}{1-c}$ the following are equivalent

$$(3) \quad (i) \frac{E(X_n)}{n} \rightarrow c; \quad (ii) \frac{E(X_n^\mu)}{n^\mu} \rightarrow \frac{c}{\mu + c(1-\mu)} \quad (n \rightarrow \infty).$$

Proof. Denote $C(y) := \sum_{k \leq y} p_k$. Then ABEL's partial summation gives

$$(4) \quad \sum_{k \leq y} k^\sigma p_k = y^\sigma C(y) - \sigma \int_1^y t^{\sigma-1} C(t) dt, \quad \sigma \in \mathbb{R}.$$

Putting $n = [y]$ in $\lim_n \frac{1}{n} E(X_n) = c$ and $\sigma = 1$ in (4), we obtain:

$$\int_1^y C(t) dt / (yC(y)) \rightarrow 1 - c \quad (y \rightarrow \infty).$$

Therefore, Lemma 3 gives $C(y) \in K_{c/(1-c)}$.

Applying again this lemma with $\rho = \frac{c}{1-c}$, $\sigma = \mu$, from (4) we get that

$$\frac{\sum_{k \leq y} k^\mu p_k}{y^\mu C(y)} = 1 - \mu \frac{\int_1^y t^{\mu-1} C(t) dt}{y^\mu C(y)} \rightarrow 1 - \mu \frac{1}{\mu + \frac{c}{1-c}} = \frac{c}{\mu + c(1-\mu)} \quad (y \rightarrow \infty),$$

which gives the right-hand side of (3) with $y = n$.

Conversely, supposing the validity of

$$\frac{1}{n^\mu} E(X_n^\mu) \rightarrow \frac{c}{\mu + c(1-\mu)} \quad (n \rightarrow \infty)$$

with $n = [y]$, then (4) gives

$$y^\mu C(y) / \int_1^y t^{\mu-1} C(t) dt \rightarrow \mu + \frac{c}{1-c} \quad (y \rightarrow \infty)$$

i.e. by Lemma 3, $C(y) \in K_{c/(1-c)}$. Using again Lemma 3. with $\sigma = 1$, $f := C$ we obtain the left-hand side of (3).

The case $c = 1$ needs a different approach.

Proposition 2. For each $\mu \in \mathbb{R}$, the following are equivalent

$$(i) \quad E(X_n) \sim n; \quad (ii) \quad E(X_n^\mu) \sim n^\mu \quad (n \rightarrow \infty).$$

Proof. The condition $\frac{1}{n} E(X_n) \rightarrow 1$ ($n \rightarrow \infty$) implies

$$(5) \quad \int_1^y C(t) dt / (yC(y)) \rightarrow 0 \quad (y \rightarrow \infty).$$

Denote by A the class of functions $C(y)$ satisfying (5). We have

Proposition 2.1. The following are equivalent

$$(i) \quad C(y) \in A; \quad (ii) \quad y^\sigma C(y) \in A,$$

for each $\sigma \in \mathbb{R}$.

For the proof we need two lemmas below.

Lemma 2.1. If $C(y) \in A$ then $y^\sigma C(y) \rightarrow \infty$, ($y \rightarrow \infty$) for any fixed $\sigma \in \mathbb{R}$.

Proof. Since $yC(y) / \int_1^y C(t) dt \rightarrow \infty$ ($y \rightarrow \infty$), we can find $y_0 > 1$ such that

$$(6) \quad yC(y) / \int_1^y C(t) dt > |\sigma| + 2, \quad y > y_0,$$

i.e.

$$(7) \quad D \left(\log \int_1^y C(t) dt \right) > \frac{|\sigma| + 2}{y}, \quad y > y_0.$$

Integrating (7) over $[y_0, y]$, we obtain

$$\int_1^y C(t) dt > c(y_0, |\sigma|) y^{|\sigma|+2}, \quad y > y_0$$

i.e. taking in account (6),

$$y^\sigma C(y) > c'(y_0, |\sigma|) y^{\sigma+|\sigma|+1}, \quad y > y_0,$$

and the assertion follows.

Lemma 2.2. If $C(y) \in A$, then $\int_1^y t^\sigma C(t) dt \sim y^\sigma \int_1^y C(t) dt$ ($y \rightarrow \infty$) for any fixed $\sigma \in \mathbb{R}$.

Proof. According to Lemma 2.1 and (5), we have

$$\frac{\int_1^y t^\sigma C(t) dt}{y^\sigma \int_1^y C(t) dt} \rightarrow \frac{D \left(\int_1^y t^\sigma C(t) dt \right)}{D \left(y^\sigma \int_1^y C(t) dt \right)} = \frac{1}{1 + \sigma \frac{\int_1^y C(t) dt}{yC(y)}} \rightarrow 1 \quad (y \rightarrow \infty),$$

where D denotes d/dy .

Proof of Proposition 2.1.

If $C(y) \in A$ then, according to Lemma 2.2,

$$\frac{\int_1^y t^\sigma C(t) dt}{y^{\sigma+1} C(y)} = \frac{\int_1^y C(t) dt}{y C(y)} \frac{\int_1^y t^\sigma C(t) dt}{y^\sigma \int_1^y C(t) dt} \rightarrow 0 \quad (y \rightarrow \infty),$$

i.e. $y^\sigma C(y) \in A, \forall \sigma \in \mathbb{R}$.

Conversely, if for some $\sigma \in \mathbb{R}, y^\sigma C(y) \in A$ then

$$\frac{\int_1^y C(t) dt}{y C(y)} = \frac{\int_1^y t^{-\sigma} (t^\sigma C(t) dt)}{y C(y)} \sim \frac{y^{-\sigma} \int_1^y t^\sigma C(t) dt}{y C(y)} = \frac{\int_1^y t^\sigma C(t) dt}{y^{\sigma+1} C(y)} \rightarrow 0 \quad (y \rightarrow \infty),$$

i.e. $C(y) \in A$.

Using Proposition 2.1 and the fact

$$\frac{\sum_{k \leq y} k^\sigma p_k}{y^\sigma C(y)} = 1 - \sigma \frac{\int_1^y t^{\sigma-1} C(t) dt}{y^\sigma C(y)}$$

with $\sigma = 1$ and $\sigma = \mu$, the proof of the Proposition 2 readily follows. \square

Propositions 1 and 2 are basic for the estimation of $E(K_\mu(X_n))$. Namely, putting in Lemma 4,

$$q_{nk} := \frac{\mu + c(1 - \mu)}{cC(n)} \left(\frac{k}{n}\right)^\mu p_k, \quad 0 < c < 1, \delta > 0, \mu > \delta - \frac{c}{1 - c},$$

Proposition 1 gives $\sum_{k \leq n} q_{nk} \rightarrow 1 \quad (n \rightarrow \infty)$, and the condition from Lemma 4 is satisfied with $\varepsilon = \delta/2$. Thus we have

Proposition 3. For $0 < c < 1, \delta > 0, \mu > \delta - \frac{c}{1 - c}$, $E(X_n) \sim cn \quad (n \rightarrow \infty)$, implies

$$E(X_n^\mu \ell(X_n)) \sim \frac{c}{\mu + c(1 - \mu)} n^\mu \ell(n) \quad (n \rightarrow \infty).$$

In the same way, using the Proposition 2, we get

Proposition 4. If $E(X_n) \sim n \quad (n \rightarrow \infty)$, then

$$E(X_n^\mu \ell(X_n)) \sim n^\mu \ell(n) \quad (n \rightarrow \infty),$$

for any $\mu \in \mathbb{R}$.

In the case $c = 0$, Lemma 3 gives $C(y) \in K_0$ and we have

$$(8) \quad \forall \mu > 0 : E(X_n^\mu) = o(n^\mu) \quad (n \rightarrow \infty).$$

Hence,

Proposition 5. *If $\frac{1}{n} E(X_n) \rightarrow 0$ ($n \rightarrow \infty$), then*

$$E(X_n^\mu \ell(X_n)) = o(n^\mu \ell(n)) \quad (n \rightarrow \infty),$$

for any $\mu \in \mathbb{R}^+$, $\ell \in K_0$.

Proof. Applying Lemma 2 and (8), we get

$$\begin{aligned} \frac{E(X_n^\mu \ell(X_n))}{n^\mu \ell(n)} &= O\left(\sup_{k \leq n} (k^{\mu/2} \ell(k))\right) \frac{E(X_n^{\mu/2})}{n^{\mu/2} \ell(n)} \\ &= O(n^{\mu/2} \ell(n)) \frac{o(n^{\mu/2})}{n^{\mu/2} \ell(n)} = o(1) \quad (n \rightarrow \infty). \end{aligned}$$

It is evident that, summarising results from Propositions 3, 4 and 5, for the generalised moments of positive order we can formulate the following

Proposition 6. *If $\frac{1}{n} E(X_n) \rightarrow c$, ($n \rightarrow \infty$), $c > 0$, then*

$$E(X_n^\mu \ell(X_n)) \sim \frac{c^{1-\mu}}{\mu + c(1-\mu)} (E(X_n))^\mu \ell(E(X_n)), \quad \mu \in \mathbb{R}^+, \ell \in K_0 \quad (n \rightarrow \infty);$$

therefore giving an answer to the question posed in (2) in the case of the sequential law.

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